A Fast Algorithm for Undetermined Mixing Matrix Identification Based on Mixture of Guassian (MoG) Sources Model

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Abstract—This paper proposes a new fast method for identifying the mixing matrix based on a binary state mixture of Gaussian (MoG) source model. First, a necessary discussion for solving the mixing matrix detection is offered under the multiple dominant circumstance. Second, a density detection method is proposed to improve the identification performance. Simulations are given to demonstrate the effectiveness of our proposed approach.

Index Terms—blind source separation (BSS); sparse component analysis (SCA); density detection; mixture of Gaussian (MoG) model.

I. INTRODUCTION

As a result of the widely application in the area of speech recognition, wireless communication and biological medical signal processing, Blind Source Separation (BSS) [1-6] is becoming one of the hottest spots in the signal processing field. The linear model of BSS can be stated as follows:

\[ x(t) = As(t) + e(t) \]  

(1)

where \( t = 1, 2, \ldots, T \) and \( A = [\alpha_1 \ldots \alpha_n] \in R^{m \times n} \) is the mixing matrix. The sample of sources \( s(t) \in R^{n \times 1} \), the observed signals sample \( x(t) \in R^{m \times 1} \) and the white Gaussian noise sample \( e(t) \in R^{m \times 1} \). If \( m < n \), which means that the number of sources is greater than the number of observed signals, then the separation problem is degenerated to the Underdetermined Blind Source Separation (UBSS) [7] problem. In this case, traditional independent component analysis (ICA) cannot be directly applied again. However, since signals are sparse in the real environment or in the frequency domain through Fourier or Wavelet transform, therefore we can solve the UBSS problem using sparse component analysis (SCA) [8]. According to the sparsity assumption of sources, model (1) can rewrite as follows:

\[ x(t) = \sum_{j=1}^{k} \alpha_{s_j}(t)s_{s_j}(t) + e(t), i_j(t) \in \{1, 2, \ldots, n\} \]  

(2)

where \( t = 1, 2, \ldots, T \) and \( k \) is the number of active source components at each instant.

Typically, a two-stage “clustering-then-\( l_1 \)-optimization” approach is often used in SCA, which is included by the mixing matrix estimation stage and sources recovery stage. According to the difference of active components number of sources, the problem of mixing matrix estimation can be categorized into two types: \( k = 1 \) and \( k > 1 \). To the one single dominant SCA problem [7] \((k = 1)\), several linear orientation-based algorithms [7-9] are addressed to solve this single dominant SCA problem; To the second case \( k > 1 \), this problem, which is called as multiple dominant SCA problem [10-12], can be solved by two steps: concentration hyperplanes clustering and mixing vectors identification. Although these hyperplane methods can effectively improve the precision of mixing matrix identification, they may not be applied in practice because of high computational costs. For overcoming this problem, we propose a novel method to reduce the time cost in the identification of \( A \). First, we discuss the geometrical distribution feature of the observed sample. Second, we give a simple density detection method to reduce the complexity of algorithm by avoiding traditional concentration hyperplane clustering step.

This paper is organized as follows: we make an intuitive and heuristic analysis of new algorithms in section 2. The binary state mixture of Gaussians (MoG) mode [13] is introduced in section 3. The complete algorithm is given in section 4. Section 5 provides some numerical simulations to demonstrate the effectiveness of our new algorithm. Then we discuss and conclude in section 6.

II. THEORETICAL ANALYSIS OF THE ALGORITHM

A. Evolutionary Algorithm

From the k-sparse mixture model of (2), there are \( c(e = C_k^n) \) \( k \)-dimensional hyperplanes to the observed data samples. And each column of lies in the hyperplanes.
of \( q(q = C_{n-1}^{k-1}) \) hyperplanes. By the normalization process for each \( x(t) \), data samples are projected onto a unit \( n \)-dimensional sphere (as is shown in Fig. 1). In other words, the directions of vectors in mixing matrix \( A \) can be detected by the directions of data sample intersections. Therefore, we can estimate the columns of mixing matrix \( A \) by detecting these intersections instead of hyperplane clustering.

Since the amplitude of sources is limited, then each hyperplane is bounded in a fixed region in which the area is considered as \( s \). Suppose the probability of data points located in a hyperplane is \( f_x \) and one hyperplane can be equally divided into several hyperplanes; Without loss of generality, the probability of data points locating in a hyperplane, which is denoted as \( \phi \), can be calculated as \( \phi = \int f_x dx_2 dx_3 ... dx_m \). For simplicity, we assume that each hyperplane has the same area. If the number of hyperplanes is \( l \) and the number hyperplanes is denoted hyperplanes are denoted as \( N_1, ..., N_c \), hyperplaneprobability of points in the same hyperplane (which is denoted as \( \phi \)) is approximately as \( f_x \times \frac{s}{l} \). As is stated in section 2.1, there are \( n \) intersections on the hypersphere. The probability of the hyperplane containing intersection point is denoted as \( p_i (i \in \{1...n\}) \), which is intersected by \( q \) hyperplanes \( B_i_1, ..., B_i_q \). Therefore, the number of points in this hyperplane can be calculated by the following equation:

\[
N_{p_i} = (N_{B_i_1} \times f_{B_i_1}(P_i) + ... N_{B_i_q} \times f_{B_i_q}(P_i)) \times s/l \tag{3}
\]

We assume that all the hyperplanes have the same number of points, and this total number equals to \( N \). Then (3) can be changed as

\[
N_{p_i} = (f_{B_i_1}(P_i) + ... f_{B_i_q}(P_i)) \times N \times s/l \tag{4}
\]

But the other hyperplanes which do not contain intersection points would contain less than \( N \times f_{B_i}(P_i) \times s/l \) numbers of point, \( i \in \{1...c\} \), where \( f_{B_i}(P_i) \) refers to probability density functions of other points in an arbitrary hyperplane \( B_i \), \( i \in \{1...c\} \). In other words, the number of points in the two kinds of hyperplane (one contains the intersection point and the other does not) is greatly different. So the ratio between them is given as:

\[
\frac{N_{p_i}}{N_{p_i}^*} = \frac{f_{B_i_1}(P_i) + ... f_{B_i_q}(P_i)}{f_{B_i}(P_i)} \tag{5}
\]

Note that if the distribution of points in every hyperplane is previously known, there may be some methods to distinguish the difference between intersection regions and other regions. For example, consider points in all hyperplanes are identically distributed with a uniform distribution. Then the ratio value of these two kind regions is

\[
\frac{N_{p_i}}{N_{p_i}^*} = \frac{f_{B_i_1}(P_i) + ... f_{B_i_q}(P_i)}{f_{B_i}(P_i)} = q \tag{6}
\]

As is shown in (6), the number of data points in the intersection regions is \( q \) times larger than other regions which do not contain intersection pointintersection detect the intedenss points from the density regions.

III. SYSTEM MODEL AND THE DISTRIBUTION FEATURE OF DATA SAMPLES

From the analysis above, we found that the distribution of observed data points in \( n \)-dimensional space is decided by the distribution of observed signal points in one hyperplane. In this section, our major job is to study the distribution of observed signals in a hyperplane.

A. The mathematical model of one hyperplane

Still review the model of (2), suppose that there are \( N \) ohyperplanes in the same hyperplane. Then the system model is:

\[
x_{t_k} = \sum_{l=1}^{k} \alpha_{i_l} s_{i_l}(t_{i_l}) + e(t_{i_l}), l \in 1...N, i_j \in 1, 2, ..., n \tag{7}
\]

As is stated in (7), it is very important to study the source model. In order to depict the distribution of source, we will introduce the following source models.

B. Binary state MoG source model

The mixture of Gaussian model [13] is one of the non-Gaussian signal model. A \( p \)-th order of MoG is given as:

\[
p(s_i) = \sum_{k=1}^{p} \pi_{i,k} N(s_i, 0, \sigma^2_{i,k}) \tag{8}
\]

where \( \sum_{k=1}^{p} \pi_{i,k} = 1 \). The MoG model is often used for depicting non-Gaussian signals like speech/audio signals. To the binary state of MoG model, which is the simplest MoG model, is widely applied in image processing and modeling sparsity in wavelet decomposition [13-16]. This model is provided as follows:

\[
p(s_i) = \pi_{i,1} N(s_i, 0, \sigma^2_{i,1}) + \pi_{i,2} N(s_i, 0, \sigma^2_{i,2}) \tag{9}
\]

where \( \delta_{1,1} \approx 0, \delta_{1,2} > 0, \pi_{i,1} \in [0, 1] \) and \( \pi_{i,1} + \pi_{i,2} = 1 \). In other words, we can rewrite model (9) as follows:

\[
p_{s_i} = \begin{cases} 
N(s_i, 0, \sigma^2_{i,1}), & s_i \text{ is inactive with probability of } \pi_{i,1} \\
N(s_i, 0, \sigma^2_{i,2}), & s_i \text{ is inactive with probability of } \pi_{i,2} 
\end{cases} \tag{10}
\]
C. Distribution of data points in one hyperplane

Suppose the components of source signals are independent with the same distribution, then the distribution of observed signals is:

\[
P(x|A, s) = \sum_{i=1}^{k} N_k(0, \alpha_i^T \delta_2^2) = N_k(0, \sum_{i=1}^{k} \alpha_i^T \delta_2^2)
\]

Therefore, we can know that the distribution of the observed signals also satisfies the Gaussian distribution. The Gaussian signals, the values of probability density function in the interval \([-\delta, \delta]\) are close to each other. When the variance is large enough, especially when \(\delta \to \infty\), the probability of hypersphere points in a hypersphere can be viewed as a constant in the interval \([-\delta, \delta]\). Therefore, the value of (5) is close to \(q\) which is large enough to distinguish intersection regions from others. For example, assuming that 5 sources are generated by the model of binary state sparse MoG and the parameter \(\delta_2 = 1\), then produce 10000 observed points in a 3-dimensional space. We divide the data sample space into 350 hyperplanes and calculate the number of each hyperplane. As is shown in Fig.2, it is obvious that there are 5 indices are more significant than others.

IV. COMPLETE MIXING MATRIX IDENTIFICATION ALGORITHM

Here we summarize the complete algorithm of mixing matrix identification:

1) Remove the sample that are close to origin.
2) Normalize and symmetric the sample \(x(t)\) by the following process:

\[
x(t) = \begin{cases} 
\frac{x(t)}{||x(t)||}, & x(t) > 0 \\
\frac{-x(t)}{||x(t)||}, & x(t) < 0 
\end{cases}
\]

3) Divide the \(m\)-dimensional Euclidean space into hyperplanes, where \(l_i = \frac{\max(x(t)) - \min(x(t))}{\eta}\), \(\eta\) is an interval length.
4) Assign sample \(x(t)\) into different spaces by the following method. Define a partition matrix \(U \in R^{l \times T}\), \(u_{i,j} \in [0, 1], i \in 0...L, j \in 0...T\).

For \(j = 1 : T\)  
\[
Loc = \text{ceil}\left(\frac{x_q(j)}{\eta}\right)
\]
\[
Set = u_{i,j} + Loc,j
\]

End

5) Calculate the number of points in each hyperplane. Choose the first \(n\) largest hyperplanes and estimate the center of each hyperplane by the following equations:

\[
\bar{\alpha}_1 = \frac{\sum x_1(i)}{K_1},
\]
\[
\vdots
\]
\[
\bar{\alpha}_m = \frac{\sum x_m(i)}{K_m}
\]

Where \(K_i\) is the number of the data points in the \(i\)-th data hyperplane.
6) Finally, construct vectors \([\bar{\alpha}_1, ... \bar{\alpha}_m]\) as the estimated mixing matrix \(A\).

V. SIMULATION EXAMPLES

In all experiences, source samples are generated independently and satisfy the distribution of the binary state MoG model which is also used in paper [13]. All the simulations were performed in MATLAB7 environment using Intel Pentium 42.4GHz processor with 512M RAM under Microsoft Windows XP operating system.

A. Experiment 1

Set \(n = 5, m = 4, k = 3\), the mixing matrix \(A\) are randomly generated and normalized as follows:

\[
A = \begin{pmatrix}
0.7930 & 0.7428 & 0.1410 & 0.9021 & 0.3281 \\
0.1480 & 0.5901 & 0.7010 & -0.3691 & -0.4419 \\
-0.5910 & 0.3161 & -0.6992 & -0.2235 & -0.8349
\end{pmatrix}
\]

The Procedures of our algorithm are shown in Fig.3 and we obtained the estimated mixing matrix as follows:

\[
\tilde{A} = \begin{pmatrix}
0.7890 & 0.9030 & 0.7441 & 0.3244 & 0.1434 \\
0.1546 & -0.3665 & 0.5888 & -0.4454 & 0.7000 \\
-0.5941 & -0.2231 & 0.3152 & -0.8343 & -0.6995
\end{pmatrix}
\]

For demonstrating the validity of our algorithm, the criterion which is presented in paper [11] and paper [12] is used:

\[
\xi = \min_{P \in P} ||A - \tilde{A}P||_2
\]

Where \(\rho\) is the set of all permutation matrices. We calculate estimation error is 0.0089, and the result is 0.0066 using the algorithm in paper [13] and 0.2018 with paper [17]. The process took about 160s when the source sample
number is 1000 and 25s when the source sample number is 400, while the algorithm in this paper only took less than 1s to finish the computing process.

B. Experiment 2

In order to demonstrate our fast detection algorithm is capable of solving large scale problems, we compared our algorithm to the existing method proposed in paper [13] by using two simulations. In the first simulation, the parameters are set as \( n = 15, m = 7, k = 3, T = 9000 \); in the second experiment, parameters are set as \( n = 30, m = 15, k = 2, T = 8500 \). Our method took less than 6 seconds in this two simulation while it took about 40 minutes for the first case and two hours for the second case by using the algorithm in paper [13].

<table>
<thead>
<tr>
<th>Parameters</th>
<th>Table 1.</th>
<th>Time</th>
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<tbody>
<tr>
<td>( n=15, m=7, k=3, T=9000 )</td>
<td>6 Sec</td>
<td>About 40 minutes</td>
</tr>
<tr>
<td>( n=30, m=15, k=2, T=8500 )</td>
<td>4 Sec</td>
<td>About 2 hours</td>
</tr>
</tbody>
</table>

Table 1 The performance comparison between our algorithm and the algorithm in paper [13]. To measure the precision of identification, the angle between each estimated vector and its corresponding actual mixing vector (inverse cosine of their dot product) is calculated, the result is shown in Fig.4, the accuracy is close to that presented in paper [17], which shows the proposed method can also estimate the mixing matrix successfully.

As is seen in this experiment, the proposed algorithm can estimate mixing matrix very fast with higher accuracy and do not change much when the parameters get larger. It means that the proposed algorithm can be used for dealing with middle scale problems.

VI. DISCUSSION AND CONCLUSION

In this paper, we propose a fast algorithm to estimate the mixing matrix \( A \) in multi-dominant SCA based on a binary state sparse MoG model. There are some aspects we need to discuss to our algorithm.
state mixture models have been used very successfully in the salient features for very sparse data. In fact, binary models are too simple, it has sufficient complexity to capture model. Although such a source model might be considered, which will leads a bad performance.η gets hyperplane number of hyperplane is increasing, according to the situation of reality. Generally, when η is hyperplane number of hyperplane is increasing, which will improve the identification accuracy of mixing vector. However, the existing algorithms can estimate the mixing vector precisely if the system is noise-free. But it should be mentioned that, the estimation result is imprecise or even worse when noise exists by using existing algorithms. In contrast, the proposed algorithm in this paper can smooth the noise effect and achieve a good accuracy.

A. Comparison with existing algorithms in accuracy and speed efficiency

The existing algorithms to estimate the mixing matrix is based on the hyperplanes clustering idea. There are several factors that traditional method need to take into account, such as the number of samples, observations and active source component. These algorithms would suffer from an exponential growth in the computation cost like in [13]. Furthermore, if experimental data contain noises, the algorithm would convergence much slower. But our algorithm only takes few seconds to finish this process and the computation cost doesn’t change much with the growth of sample number. As observed in the simulation results, the location of mixing vector is less precise when there is noise-free, because the proposed algorithm regards the center of the intersection point region as the estimation of the mixing vector. However, the existing algorithms can estimate the mixing vector precisely if the system is noise-free. But it should be mentioned that, the estimation result is imprecise or even worse when noise exists by using existing algorithms. In contrast, the proposed algorithm in this paper can smooth the noise effect and achieve a good accuracy.

B. Comments on choosing the parameter

In the proposed algorithm, the parameter η of hyperplane—ethe size of hyperplane should choose a suitable value according to the situation of reality. Generally, when η is hyperplane number of hyperplane is increasing, which will improve the identification accuracy of mixing matrix. And if η gets hyperplane number of hyperplane is deleading, which will leads a bad performance.

C. Comments on prior knowledge

Another previous knowledge is that the source signals should distribute in the binary state sparse distribution model. Although such a source model might be considered too simple, it has sufficient complexity to capture the salient features for very sparse data. In fact, binary state mixture models have been used very successfully by Olshausen and Millman [12] to estimate overcomplete bases for images. Fortunately, it will be possible to solve all kinds of BSS problem of different kinds of source models if the probability density function of observed data points is given in advance. These issues are currently under study. Meanwhile, the construction of the sparse signal model is still an open problem that remained to be solved.

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